

## The reporting odds ratio versus the proportional reporting ratio: ‘deuce’<sup>†</sup>

An article published in this issue by Rothman *et al.* entitled ‘The reporting odds ratio and its advantages over the proportional reporting ratio’<sup>1</sup> argues that the reporting odds ratio (RORs) is a more valid measure of association when applied to datasets of spontaneous reports of suspected adverse reactions. However, in our view, this paper fails to provide a coherent basis to support its title, conclusions and take home messages. It has brought together two different issues and confused them. The first issue is what measure is to be used to identify associations, and the second is what comparisons are to be made within a database.

RORs<sup>2</sup> and proportional reporting ratios (PRRs)<sup>3</sup> are both measures of disproportionality used for the purpose of detecting signals in spontaneous ADR reporting databases. Both are calculated from the same  $2 \times 2$  tables with the PRR being identical to the calculation of a relative risk (RR) from a cohort study i.e.  $(a/a + c)/(b/b + d)$ <sup>†</sup> and the ROR identical to the calculation of an odds ratio (OR) from a case-control study i.e.  $ad/bc$ . It is well-recognised that these measures will give very similar results providing, as is virtually always the case in this context,  $a$  is a small proportion of  $a + c$  and  $b$  is a small proportion of  $b + d$ . Effectively this is a similar argument used to show that the OR in a case-control study approximates the RR.<sup>4</sup> Whilst the calculations of PRR and RR, and ROR and OR respectively are identical, it is important to understand that when used in this context they are not meant to actually estimate RR but to assist in efficiently identifying potential drug hazards from often large datasets of spontaneous reports of suspected adverse reactions.

A judgment on the validity and utility of these measures should be based on comparison of their sensitivity, specificity and predictive values in signal detection from a real dataset. Since Rothman *et al.* make only passing reference to a paper published in this Journal in 2002,<sup>5</sup> the reader might be forgiven for assuming that such data do not yet exist. In fact, that 2002 paper<sup>5</sup> made such a comparison and showed clearly that, in practice, there is no important difference between the measures for the purpose for which they are used. However, it was pointed out that there are some minor issues which might influence choice of measure. In particular, making adjustments in a logistic regression analysis is easy with an ROR but the ROR will occasionally be impossible to calculate i.e. when  $b$  or  $c$  is zero<sup>5</sup> whereas the PRR can still be calculated when  $c$  is zero (but not when  $b$  is zero). Statistical arithmetic is unimportant here and the differences between these measures of association in this setting are of no major significance.

The second, and possibly more interesting issue is the choice of a comparison group. This has nothing whatever to do with ORs or RRs. Rothman *et al.* use a single invented example for illustrative purposes but which is of little relevance to the usual problems faced in signal detection. The authors seem to be trying to demonstrate that inclusion of the data for event B in their tables will bias the estimate of RR for event A when the drug produces a 10-fold reduction in event A. This would be unusual and we doubt that these approaches will be valuable in such a situation. What would be more relevant from their example would be to consider whether or not event B will be affected by event A (i.e. will it lead to a spurious signal being detected?). In that situation, the ROR is actually larger (and therefore more biased?) than the PRR (1.8 vs. 1.7) but still insufficient to raise a strong signal. Thus their example seems merely to illustrate the mathematical inevitability that an ROR will always be further away from 1 than the PRR. However, the precise values of the point estimates derived from such calculations are of

\*Correspondence to: P. Waller, Consultancy in Pharmacovigilance and Pharmacoepidemiology, 15 Tamella Road, Botley, Southampton, 5030 2NY, UK.

E-mail: patrick.waller@btinternet.com

<sup>†</sup>In this Editorial  $a$ ,  $b$ ,  $c$  and  $d$  are defined as per the Appendix Table of Rothman *et al.*<sup>1</sup>

little importance: the critical issues are whether or not a signal will be detected and what impact such biases might have on the efficiency of signal detection. In practice, the major problem is that frequently reported adverse effects may make it difficult to detect other signals relating to the drug in question. For example, with uveitis related to rifabutin (see Table 2 in Evans *et al.*<sup>3</sup>), 75% of UK reports for that drug were for this one reaction, whilst the proportion for all other drugs was 0.1%. This is a very extreme signal which could well lead to the PRR and ROR being less than one for the association between rifabutin and other ADRs, even when important signals were present. Furthermore it would also impact on the detection of signals of uveitis with other drugs. Unfortunately, the example used by Rothman *et al.* fails to address these points.

In our view, the suggestion by Rothman *et al.* that the ROR allows estimation of RR is misguided. Rather both RORs and PRRs allow us to define whether or not a signal is present and, if it is, give a broad indication of the strength of that signal. As far as we can judge, Rothman *et al.* seem to be arguing a conceptual advantage relating to data which may cause bias. However, any such conceptual advantage is irrelevant to these circumstances since there is no reason why, when warranted, exclusions cannot be made from the PRR calculation.<sup>2</sup> Thus, whilst we agree with Rothman *et al.* that selection of controls is an important issue, we do not accept that use of an ROR rather than a PRR provides an advantage in this respect.

In conclusion, RORs and PRRs have been shown to be similarly effective measures of disproportionality<sup>5</sup> and the contrary arguments put forward by Rothman *et al.*<sup>1</sup> seem to be of no practical consequence. Bayesian methods are also useful.<sup>6</sup> We suggest that the most important current issues in this field are whether better approaches to the analysis of sponta-

neous ADR reporting data (including control selection) can be developed, and how to prioritise and handle the large numbers of signals that should now be detected.

Patrick Waller<sup>1\*</sup>, Eugène van Puijenbroek<sup>2</sup>  
Antoine Egberts<sup>3</sup> and Stephen Evans<sup>4</sup>

<sup>1</sup>Consultant in Pharmacovigilance and  
Pharmacoepidemiology, Southampton, UK

<sup>2</sup>Netherlands Pharmacovigilance Centre Lareb  
s'Hertogenbosch, The Netherlands

<sup>3</sup>Professor of Clinical Pharmacoepidemiology  
Utrecht Institute for Pharmaceutical Sciences  
Utrecht, The Netherlands

<sup>4</sup>Professor of Pharmacoepidemiology  
London School of Hygiene and Tropical Medicine  
London, UK

## REFERENCES

1. Rothman KJ, Lanes S, Sacks ST. The reporting odds ratio and its advantages over the proportional reporting ratio. *Pharmacoepidemiol Drug Safe* 2004; **13**: 519–523.
2. Meyboom RHB, Egberts ACG, Edwards IR, Hekster YA, de Koning FHP, Gribnau FWJ. Principles of signal detection in pharmacovigilance. *Drug Safe* 1997; **16**: 355–365.
3. Evans SJW, Waller PC, Davis S. Use of proportional reporting ratios (PRRs) for signal generation from spontaneous adverse drug reaction reports. *Pharmacoepidemiol Drug Safe* 2001; **10**: 483–486.
4. Armitage P, Berry G. *Statistical Methods in Medical Research* (2nd edn). Blackwell Scientific Publications: Oxford, 1987; 456–468.
5. van Puijenbroek EP, Bate A, Leufkens HGM, Lindquist M, Orre R, Egberts ACG. A comparison of measures of disproportionality for signal detection in spontaneous reporting systems for adverse drug reactions. *Pharmacoepidemiol Drug Safe* 2002; **11**: 3–10.
6. Gould AL. Practical pharmacovigilance analysis strategies. *Pharmacoepidemiol Drug Safe* 2003; **12**: 559–574.